

# Geostatistical modeling and remotely sensed data to improve dendrometric variables prediction in *Tectona grandis* L. f. stand

## Modelización geoestadística y datos de teledetección para mejorar la predicción de variables dendrométricas en el rodal de *Tectona grandis* L. f.

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### Abstract

Detailed knowledge of teak stand structure is necessary for sustainable management plans. The integration of remote sensing variables with geostatistical modeling in teak forest stands has not been sufficiently studied and, therefore, the aim was to model the spatial distribution of teak stand variables, adding covariables. The study was carried out on 19-year-old teak stand in Brazil with 213 hectares in the initial spatial of 3 m x 3 m. Geo-referenced plots of 900 m<sup>2</sup> were allocated, and forest variables were obtained after thinning. Vegetation indices were calculated from arithmetic operations conducted between the Landsat image bands. The interpolation of forest variables was performed by the geostatistical univariate method of ordinary kriging, as well as by the multivariate method of kriging with external drift, considering the remote sensing variables as covariables. Statistical analysis of remote sensing variables shows a weak linear correlation with teak variables, which tends to make them unviable to use as covariables in geostatistical modeling. However, kriging with external drift predicts spatial patterns of forest variables with greater detail, which results in lower possible smoothing errors than those obtained by ordinary kriging and provides more accurate recommendations for localized management in teak stand. The integration of remote sensing variables in forest inventory through geostatistics is advantageous for mapping the spatial distribution of teak stand variables.

**Keywords:** Kriging with external drift, vegetation indices, teak.

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## Resumen

El conocimiento detallado de la estructura de las plantaciones de teca es necesario para los planes de gestión sostenible. La integración de variables de teledetección con la modelización geoestadística en plantaciones de teca ha sido poco estudiada y, por tanto, el objetivo consistía en modelizar la distribución espacial de las variables del rodal de teca, incorporando covariables. El estudio se realizó en plantaciones de teca de 19 años en Brasil con 213 hectáreas en el espacio inicial de 3 m x 3 m. Se asignaron parcelas georreferenciadas de 900 m<sup>2</sup> y se obtuvieron las variables forestales después del clareo. Los índices de vegetación se calcularon a partir de operaciones aritméticas ejecutadas entre las bandas de imágenes Landsat. La interpolación de las variables forestales se procedió por el método geoestadístico univariante de kriging ordinario, así como por el método multivariante de kriging con deriva externa, considerando las variables de teledetección como covariables. El análisis estadístico de las variables de teledetección muestra una discreta correlación lineal con las variables de la teca, lo que tiende a hacer inviable su uso como covariables en la modelización geoestadística. Sin embargo, el kriging con deriva externa predice los patrones espaciales de las variables forestales con mayor detalle, lo que da como resultado a posibles errores de suavización menores que los obtenidos por el kriging ordinario, y proporciona recomendaciones más precisas para la gestión localizada en plantaciones de teca. La integración de variables de teledetección en el inventario forestal mediante geoestadística es ventajosa para cartografiar la distribución espacial de las variables de los rodales de teca.

**Palabras clave:** Kriging con deriva externa, índices de vegetación, teca.

## Introduction

*Tectona grandis* L. f., popularly known as teak, is a species from the monsoon rainforest in the Asian continent, with natural distribution in India, Myanmar, Thailand, and Laos [1], [2]. Commercial teak plantations were introduced in Brazil to ensure wood availability for regional industry [3]. Teak log production represents 0.4% of the total wood produced in Brazil, with an average annual productivity of 15 m<sup>3</sup> ha<sup>-1</sup> [4]. Currently, there are 94 thousand hectares of commercial teak plantations [5] located in the Midwest and North regions.

Teak is a fast-growing species, whose wood is considered noble for its durability, it is resistant to fire, pests, and

diseases, therefore, it is widely used in luxury furniture and shipbuilding [6]. Thus, teak is an alternative to native forest species of high economic value, providing a sustainable option for the Brazilian forest-based industry and reducing the pressure to exploit tropical rainforests [7], [8].

Detailed knowledge of teak stand structure is necessary for sustainable management plans. Thus, it is important to consider the spatial relationships of forest variables in a precision forestry context [9], [10]. Geostatistics is a technique based on the theory of regionalized variables[11] that provides analysis of the spatial distribution of variables and estimation at unsampled locations for thematic maps [12], [13].

Geostatistics has been applied to manage teak stands [14], [15], since they require a particular management regime of the main forest species, such as *Eucalyptus* and *Pinus* genera. Additionally, covariables can be used to increase the accuracy at the spatial level [16], [17]. Remote sensing variables can assist in geostatistical modeling [18] due to their low acquisition cost and the correlation between reflectance values and forest variables [19], [20].

This evidence is based on information obtained by different remote sensing technologies, which have shown wide application for forest cover representation, forest structure mapping, improving the accuracy in forest estimates, and supporting forest operations [21], [22]. However, the integration of remote sensing variables, such as vegetation indices, with geostatistical modeling in forest stands has not been sufficiently studied [16], [23], [24]. For teak plantations, this gap remains unexplored and, therefore, the aim of this study was to model the spatial distribution of teak stand variables, adding covariables.

## Material and methods

### Study area

The study was carried out in 19-year-old teak stand in Mato Grosso State, Brazil, with 213 hectares in the initial spatial of 3 m x 3 m. The region's climate is tropical Aw (Köppen), with a dry season in the winter, an average rainfall of 1,300 mm year<sup>1</sup>, and an annual average temperature of 25 °C [25]. Selective thinning was carried out at six, nine, twelve, fourteen, seventeen, and nineteen years old, with an average removal of 35% of the initial planting, and 17, 11, 5, 9, 3 of the total remaining trees, respectively. Pruning was performed in the second, third, and fourth years.

For data collection, 46 geo-referenced plots of 900 m<sup>2</sup> (30 m x 30 m) were allocated. In these plots, variables were obtained after thinning: V - stand volume (m<sup>3</sup> ha<sup>-1</sup>),

$G$  - basal area ( $m^2 ha^{-1}$ ),  $\bar{d}$  - arithmetic mean of diameters at breast height (cm),  $d_g$  - quadratic mean diameter (cm),  $d_{dom}$  - dominant diameter (cm);  $\bar{h}$  - arithmetic mean of the heights (m); and  $h_{dom}$  - dominant height (m).

### Remotely sensed data and imagery processing

Landsat 8 Operation Land Imager image was acquired, in march 2018, on the Earth Explorer portal of the United States Geological Survey [26]. The image was reprojected to Universal Transverse Mercator, Datum SIRGAS 2000 - Spindle 21 South, and an atmospheric correction was applied in QGIS software [27]. The fusion of multispectral bands with panchromatic was performed to obtain a higher spatial resolution image.

Vegetation indices were calculated from arithmetic operations carried out between the image bands: Simple Ratio (SR), Normalized Difference Vegetation Index (NDVI), Enhanced Vegetation Index (EVI), Soil Adjusted Vegetation Index (SAVI), Atmospherically Resistant Vegetation Index (ARVI), Visible Atmospherically Resistant Index (VARI), and Leaf Area Index (LAI) [28]. In addition, Principal Component Analysis (PCA) and Tasseled Cap Transformation were performed. The first one aims to produce uncorrelated principal components (PC), while the second one results in an orthogonal transformation of the sensor bands that redesigns the data in three axes: Brightness, Greenness, and Wetness [29], [30].

### Geostatistical analysis

Exploratory analysis was applied to verify the relationship between forest and remote sensing variables and the presence of outliers. These analyses consisted of the Anderson-Darling's test for data normality at a 95% probability level, linear correlation analysis, and frequency histograms to identify positive-skew distributions, which can hinder geostatistical modeling due to the weight of few observations in the estimation at unsampled locations [31].

Semivariances were calculated to identify the spatial dependence of the sampling units. Spherical, Exponential, Gaussian, Pentaspheric, and Circular semivariogram models were fitted by the gstat package [32] of R software [33] to describe the spatial dependence structure. The semivariograms were determined in four spatial directions,  $0^\circ$ ,  $45^\circ$ ,  $90^\circ$ , and  $135^\circ$ , to verify anisotropy. If detected, the angle was adjusted to obtain a semivariogram model with common parameters in all directions [31]. The best fit was selected based on the smallest weighted sum of squared deviations (WSSD) and highest coefficient of determination ( $R^2$ ).

The interpolation of forest variables was carried out by the geostatistical univariate method of ordinary kriging, as well as by the multivariate method of kriging with

external drift, considering the remote sensing variables as covariates.

### Ordinary kriging

In ordinary kriging (1), the estimate at the unsampled point results from a combination of values found in the close neighborhood [31] whose average is unknown. In this method, weights ( $\lambda$ ) are influenced by the fitted semivariogram, considering the distances between the sampling points  $Z(x)$  and the point to be estimated  $\hat{Z}_{KO}(x_0)$ . Thus, higher weights are assigned to the nearest points, which minimizes the grouping effects or data redundancy [34].

$$\hat{Z}_{KO}(x_0) = \sum_{i=1}^N \lambda_i Z(x_i) \quad (1)$$

### Kriging with external drift

Kriging with external drift (2), or universal kriging, is a method that allows estimating a variable considering the information of a covariate known throughout the area [35]. Thus, considering  $Z_1(x)$  as the primary variable and  $Z_2(x)$  as the covariate, this correlation can be described through a linear relationship. This indicates that the spatial variability of the secondary variable  $Z_2(x)$  is related to local trends  $Z_1(x)$ .

$$E[Z_1(x)] = \hat{\beta}_0 + \hat{\beta}_1 Z_2(x) + \varepsilon \quad (2)$$

The kriging methods estimates were compared by leaving-one-out cross-validation [36] using the gstat package [32]. BIAS, mean absolute error (MAE), mean percentage error (MPE), absolute root mean squared error (RMSE), percentage root mean squared error (RMSE%), and Pearson's linear correlation between observed and estimated values ( $r_{xx}$ ) were also applied.

## Results and discussion

### Exploratory analysis of teak stand and remote sensing variables

The low variability of forest variables (Table 1) is indicative of data homogeneity, which may be the result of thinning applied in teak stand. Thinning tends to change the forest structure, especially selective ones, by removing trees with smaller sizes and phytosanitary and quality problems, such as stem tortuosity and bifurcation [37], [38]. Remote sensing variables with the highest coefficient of variation (cv%) values (e.g., PC = 43.5 ) were more influenced by variations in vegetation cover of the study area, while the variables with lower values (e.g., VARI = 15.5 ) were less affected.

**Table 1.** Descriptive statistical analysis of forest and remote sensing variables.**Cuadro 1.** Análisis estadístico descriptivo de las variables forestales y de teledetección.

Variable	Minimum	Average	Maximum	Standard Deviation	cv%	Anderson-Darling
V ( $m^3 ha^{-1}$ )	162.00	394.29	536.64	81.66	20.7	1.33*
G ( $m^2 ha^{-1}$ )	21.35	39.93	50.66	6.11	15.3	1.37*
$\bar{d}$ (cm)	26.49	33.10	38.29	2.78	8.4	0.63ns
$d_g$ (cm)	26.23	34.80	41.62	3.26	9.4	0.45ns
$d_{dom}$ (cm)	32.95	39.59	48.46	3.62	9.2	0.15ns
$\bar{h}$ (m)	19.60	25.75	29.10	2.39	9.3	2.60*
$h_{dom}$ (m)	20.84	27.73	31.78	2.40	8.7	0.79*
SR	0.00	2.18	5.70	0.52	23.9	24.73*
NDVI	-1.00	0.35	0.70	0.13	36.4	258.55*
EVI	-0.47	0.28	0.69	0.11	39.0	70.07*
SAVI	-1.48	0.52	1.04	0.19	36.4	256.64*
ARVI	-1.00	0.66	1.43	0.21	31.6	192.19*
VARI	-0.10	0.79	1.29	0.12	15.5	144.50*
IAF	-1.28	0.83	2.23	0.32	38.6	73.66*
PC	13	61.76	336	26.84	43.5	975.57*
Brightness	12.72	56.48	303.16	23.94	42.4	970.32*
Greenness	-95.35	-2.72	13.69	6.16	-226.2	844.42*
Wetness	-67.81	-3.32	10.19	6.16	-185.5	491.70*

ns normal distribution according to Anderson-Darling's test (pvalue > 0.05), and \* non-normal distribution (pvalue ≤ 0.05).

Data normality was verified for variables arithmetic mean of diameters at breast height ( $\bar{d}$ ), quadratic mean diameter ( $d_g$ ), and dominant diameter ( $d_{dom}$ ) through Anderson-Darling's test (Table 1). The absence of normality was observed for remote sensing variables at a 95 probability level. Log-transformation was applied for forest variables, aiming to reduce the asymmetry of their distributions [39]. However, it was not possible to transform remote sensing variables since some vegetation indices showed negative values.

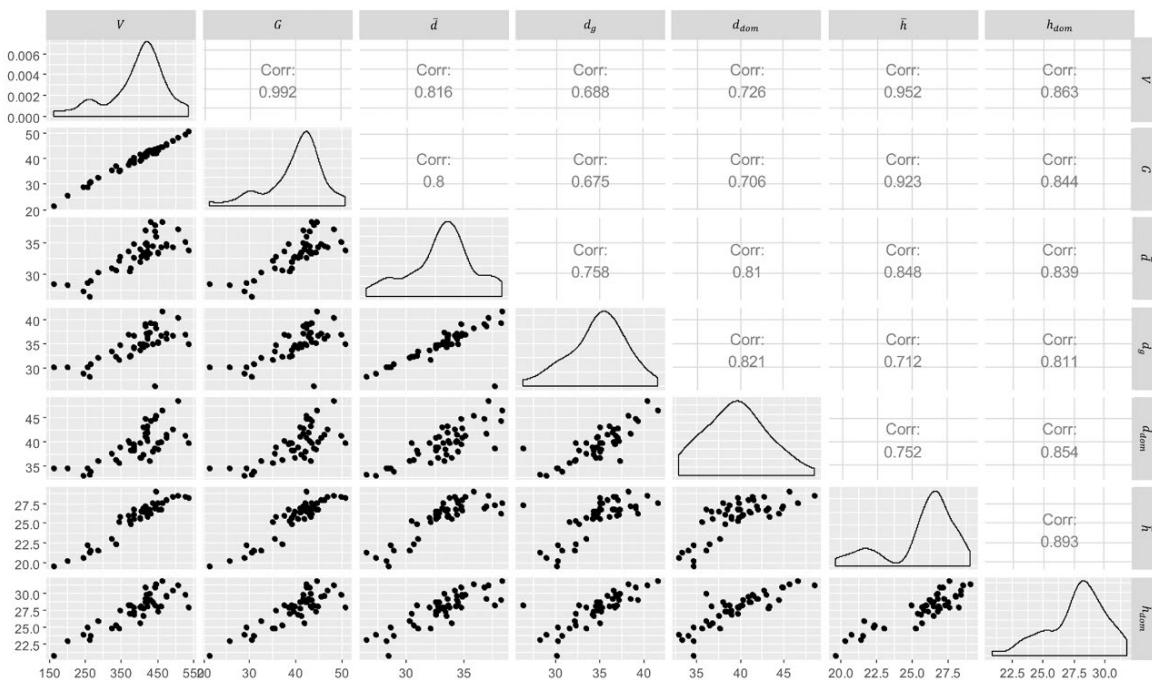
The negative-skew distributions were observed for variables volume (V) and basal area (G), in which the highest frequency class (mode) has moved to the right of the distributions (Figure 1). This result is evidence of changes in forest structure caused by thinning in long rotation teak stands [40].

Arithmetic means variables of heights ( $\bar{h}$ ) and dominant height ( $h_{dom}$ ) distributions showed similar behavior (Figure 1) since the stand structure was simplified by

successive thinning to trees with greater dominance. The tendency of bimodal distributions for these variables can indicate two site classes with different productivity. Additionally, data dispersion and correlations showed linear dependence between the forest variables. Also, a higher correlation between volume (V) and basal area (G) showed the influence of diameter in the quadratic form to determine the volume [41], [42].

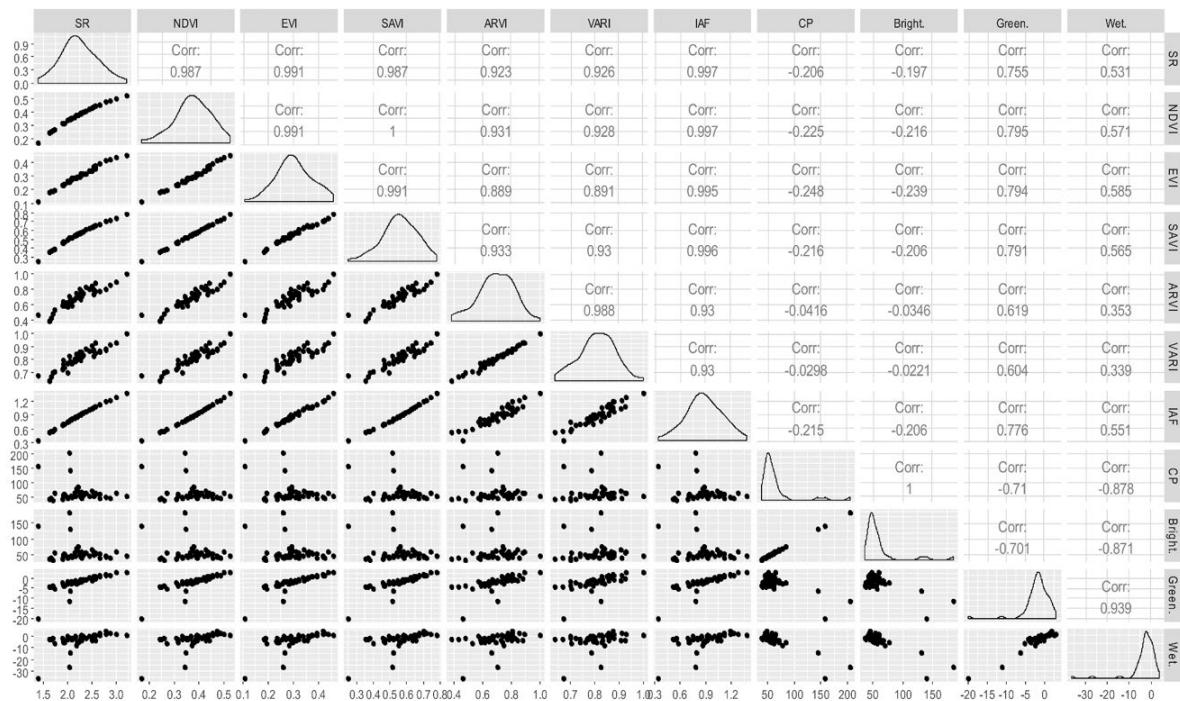
Evaluating the scatterplots of remote sensing variables (Figure 2), unimodal distributions with asymmetry were observed for Simple Ratio (SR), Normalized Difference Vegetation Index (NDVI), Enhanced Vegetation Index (EVI), Soil Adjusted Vegetation Index (SAVI), Atmospherically Resistant Vegetation Index (ARVI), Visible Atmospherically Resistant Index (VARI), and Leaf Area Index (LAI).

Moreover, the strong linear dependence between these variables is evidenced by the lower data dispersion [43], since they are derived from the same information, as



**Figure 1.** Relationships between teak stand variables.

**Figura 1.** Relaciones entre las variables de rodales de teca.



**Figure 2.** Relationship between forest and remote sensing variables.

**Figura 2.** Relación entre variables forestales y de teledetección.

the near-infrared (NIR) and red band (RED) reflectance. Principal Components (PC) and Brightness showed a strong linear dependence and positive-skew distributions, while Greenness and Wetness presented

high correlation and negative-skew distributions (Figure 2). These variables also showed extreme values, indicating possible outliers [44].

Using thematic maps of vegetation indices (Figure 3), green-colored areas indicate higher values and stronger vegetative vigor, while areas with red, orange, and yellow colors showed less vegetative. Thus, it was possible to distinguish sites with dense vegetation from areas with sparse vegetation and exposed soil (Figure 3). This behavior was inverted for PC (Figure 3g) and Brightness variables (Figure 3h).

The central linear strip in the stand (Figure 3) represented exposed soil due to a firebreak and therefore low vegetation index values were obtained. In addition, six selective thinning have been carried out since the stand's establishment, with a decrease in initial average density from 1,111 trees  $\text{ha}^{-1}$  to 139 trees  $\text{ha}^{-1}$ . The final density ranged from 89 trees  $\text{ha}^{-1}$  to 189 trees  $\text{ha}^{-1}$ , in which the lowest values of remote sensing variables also indicated the sites with less dense vegetation [45] due to more intense thinning in teak stand.

### **Relationship between forest and remote sensing variables.**

The teak stand variables showed weak linear correlations [46] with remote sensing variables (Figure 4). Higher negative values were observed between V and h with PC and Brightness, while the greatest positive correlation occurred between V with Greenness. For dg, there was not a correlation between PC and Brightness. These results may be indicating a rejection of the hypothesis formulated in the present study, in which remote sensing as covariables in geostatistical modeling does not allow predicting teak stand variables. However, spatial dependence also needs to be assessed for a complete evaluation of the data [47].

The lack of high correlations between forest and remote sensing variables can be explained by the vegetation cover aspect generated by a satellite image, which is the result of a process involving many environmental parameters and factors [48]. Therefore, a measurement obtained by a remote sensor cannot be only explained by the intrinsic vegetation characteristics since an image is interfered by the atmosphere effects, canopy attributes, radiation sources, and soils [49], [50].

### **Ordinary Kriging**

The fitted semivariograms for forest variables showed spatial dependence and better statistics with the Spherical model, except for ddom with the Exponential model. The nugget effect showed values lower than 0.02 and indicates the statistical quality of the fitted models. In addition, the range presented minimum values of 222 m for ddom and a maximum of 913 m for d. The other fits resulted in ranges between 761 and 881 m (Figure 5).

In addition, semivariances showed low dispersion around the estimated average line for forest variables

(Figure 5). The increasing behavior of the semivariance line and the subsequent stabilization around a maximum distance, defined by the range parameter, confirms the quality of semivariogram fits in this study [51]. Thus, the interpolation was performed by ordinary kriging, in which thematic maps with heterogeneous spatial patterns were obtained (Figure 6).

The central region of the maps showed the highest values of volume (V), basal area (G), mean quadratic diameter ( $d_g$ ), mean height (h), and dominant height (hdom); while the East region presented the lowest values for all forest variables (Figure 6). These results indicate the influence of site quality, which is better in areas where the forest variables had higher values [34], which can be confirmed by the greater dominant heights present in the same region (Figure 7g).

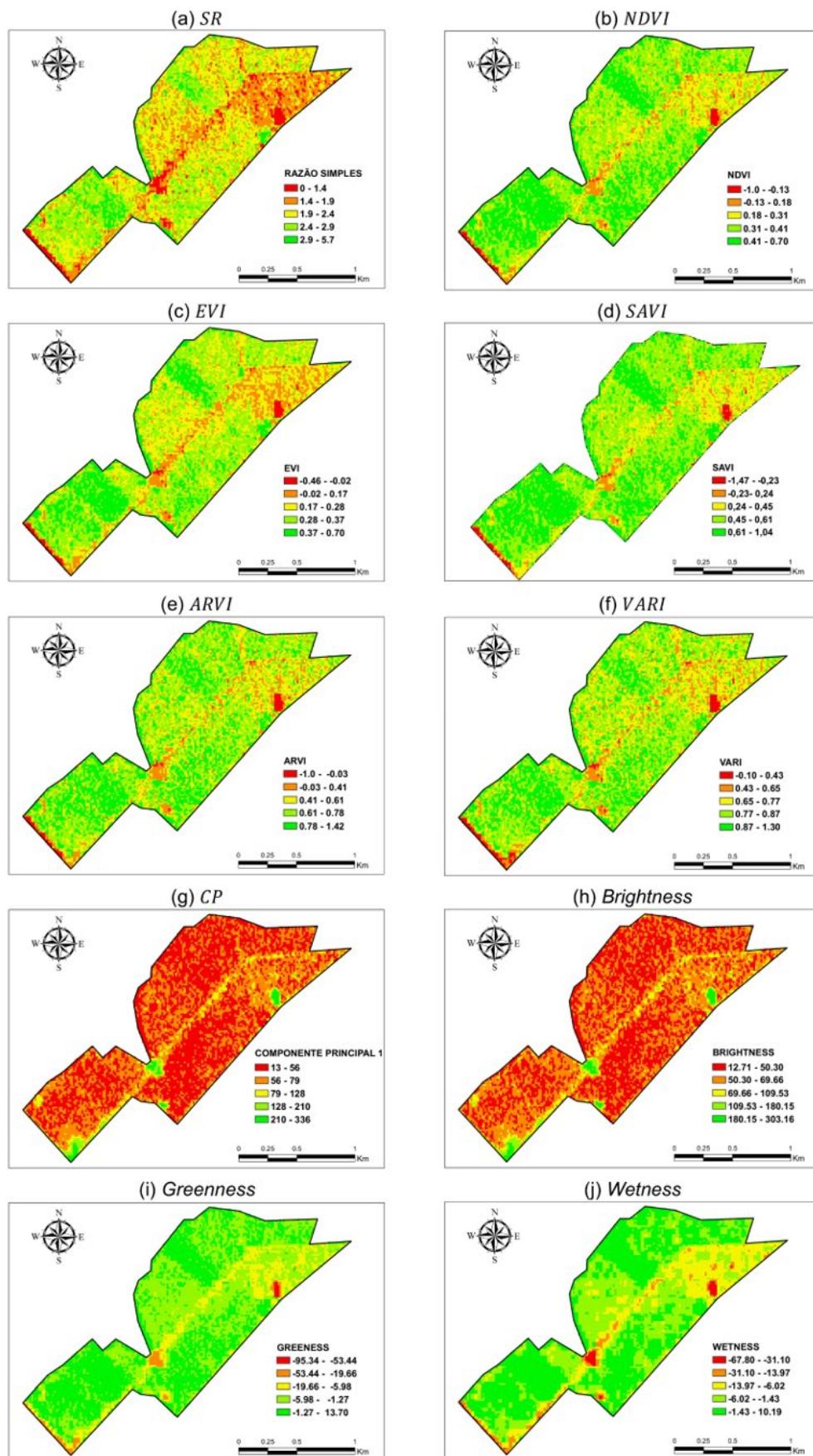
### **Kriging with external drift**

Kriging with external drift was initiated by fitting semivariograms (Figure 7), however, in this method, remote sensing covariables were used to explain a portion of forest spatial variability. The spherical model provided the best fit, except for  $d_{\text{dom}}$  with the exponential model. These fits resulted in low nugget effect values, showing the goodness of statistical quality of the modeling [52].

Range values were similar to those obtained by ordinary kriging, in which lower values were obtained for volume (V), basal area (G), arithmetic mean of the heights (h), and dominant height ( $h_{\text{dom}}$ ) with the covariable Wetness (Figure 7). These results demonstrate the influence of this covariable on the spatial dependence of forest stands [40]. On the other hand, Brightness, which expressed an index of soil brightness [30], resulted in the highest range values (Figure 7).

Covariable Greenness resulted in the best-fit volume variable (V), with the highest coefficient of determination ( $R^2$ ), equal to 0.990. Basal area (G) and arithmetic mean of heights (h) performed better with the VARI index, resulting in the smallest weighted sum of squared deviations (WSSD) of  $3.79 \times 10^{-9}$ . Covariable Wetness presented the highest  $R^2$  (0.952) and the lowest WSSD ( $5.94 \times 10^{-10}$ ) for the arithmetic mean variable of the diameters (d). The variable quadratic mean diameter ( $d_g$ ) performed better with the EVI index, while covariable Brightness enabled the highest  $R^2$  (0.965 and 0.992) and lowest WSSD ( $3.93 \times 10^{-10}$  and  $1.16 \times 10^{-10}$ ) for dominant diameter ( $d_{\text{dom}}$ ) and dominant height ( $h_{\text{dom}}$ ) (Figure 7).

Semivariograms fitted according to remote sensing covariables with the best statistical performance in kriging with external drift showed appropriate patterns, with an increasing and subsequently stabilizing behavior of the estimated semivariances (Figure 7). As a result, they were used to compose thematic maps of forest variables (Figure 8).

**Figure 3.** Thematic maps of remote sensing variables in teak plantations.**Figura 3.** Mapas temáticos de variables de teledetección en plantaciones de teca.

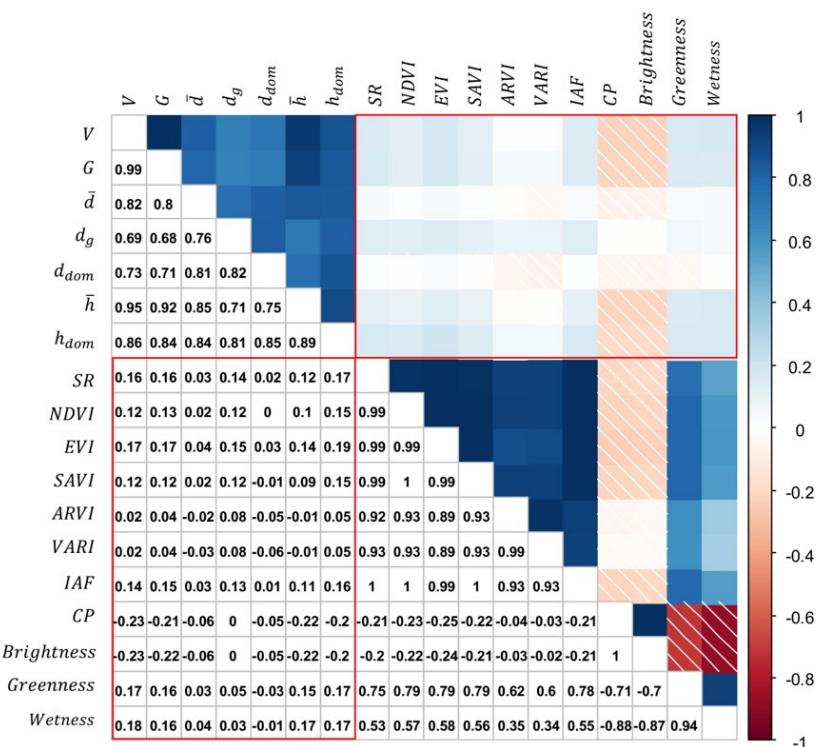


Figure 4. Correlation between teak stand variables and remote sensing covariables.

Figura 4. Correlación entre las variables del rodal de teca y las covariables de teledetección.

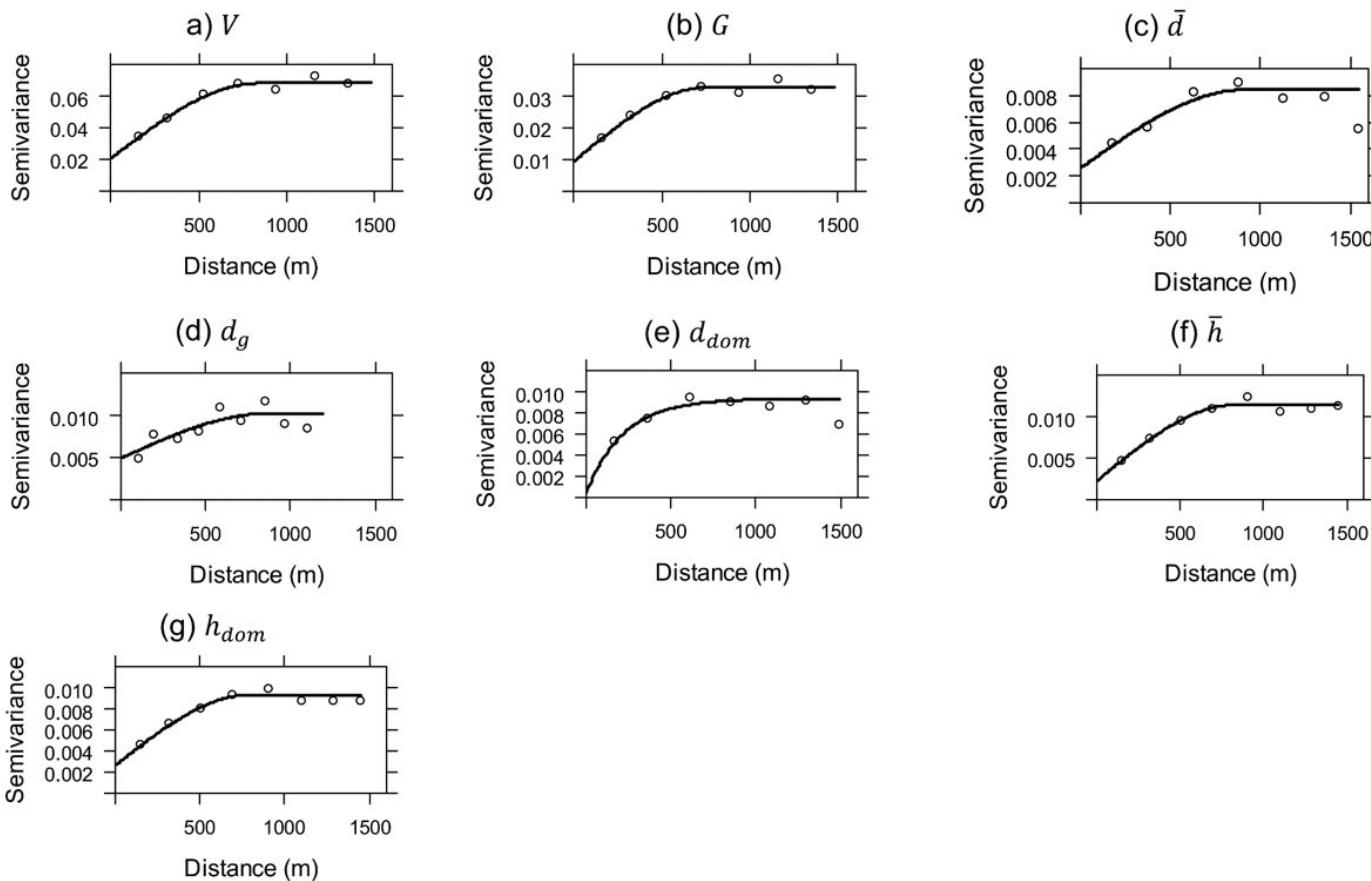


Figure 5. Semivariograms fitted for teak variables.

Figura 5. Semivariogramas ajustados para las variables de la teca.

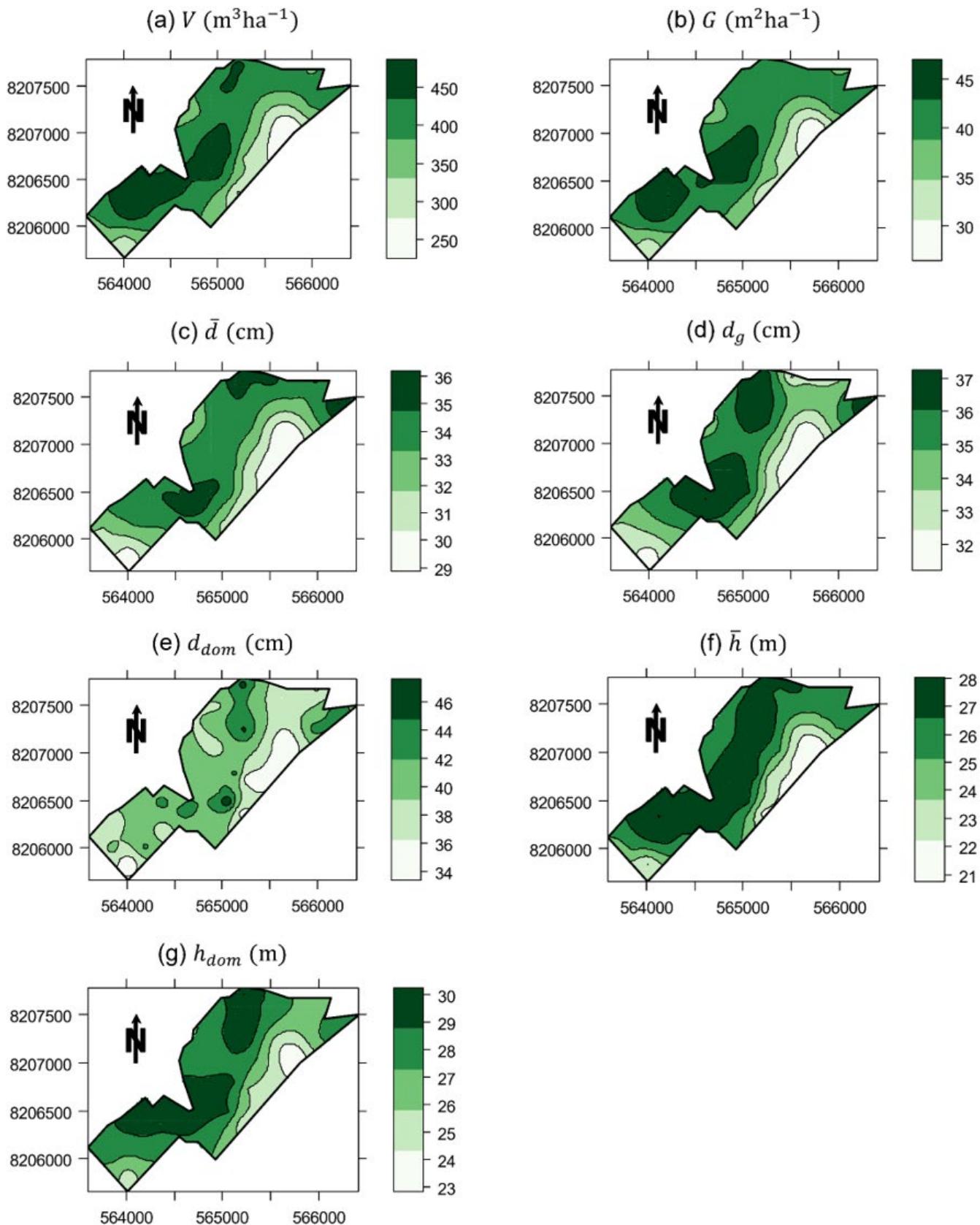


Figure 6. Spatial distribution of teak stands variables by ordinary kriging.

Figura 6. Distribución espacial de las variables de los rodales de teca mediante kriging ordinario.

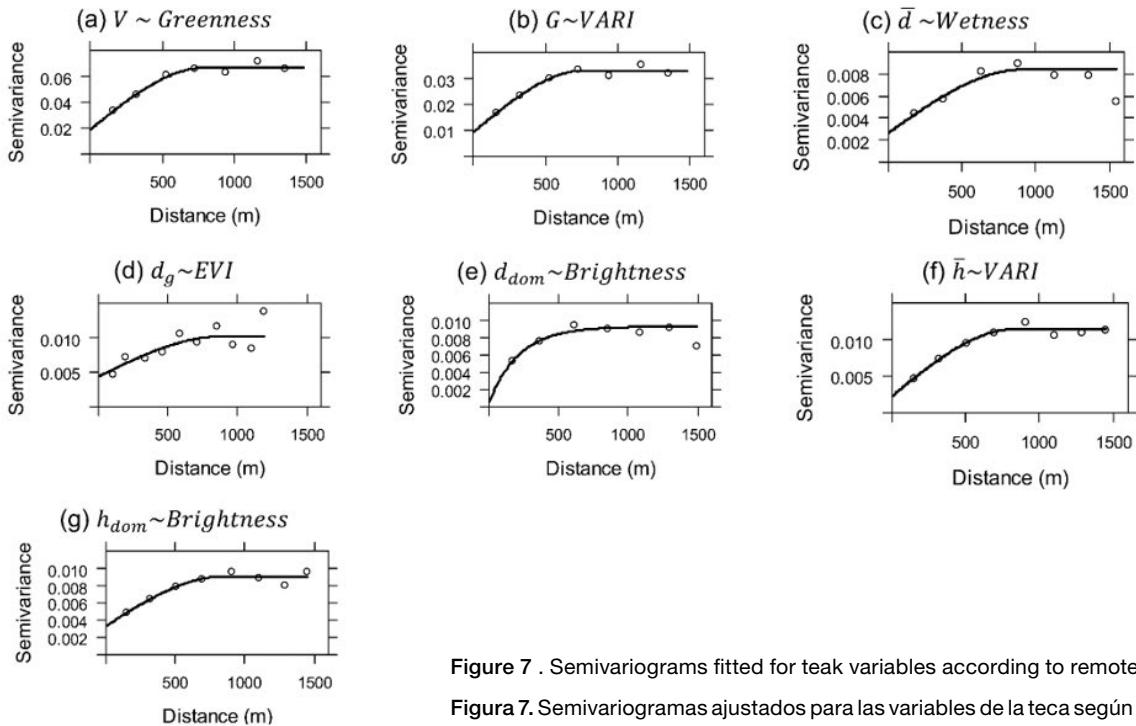


Figure 7 . Semivariograms fitted for teak variables according to remote sensing covariables.

Figura 7. Semivariogramas ajustados para las variables de la teca según las covariables de teledetección

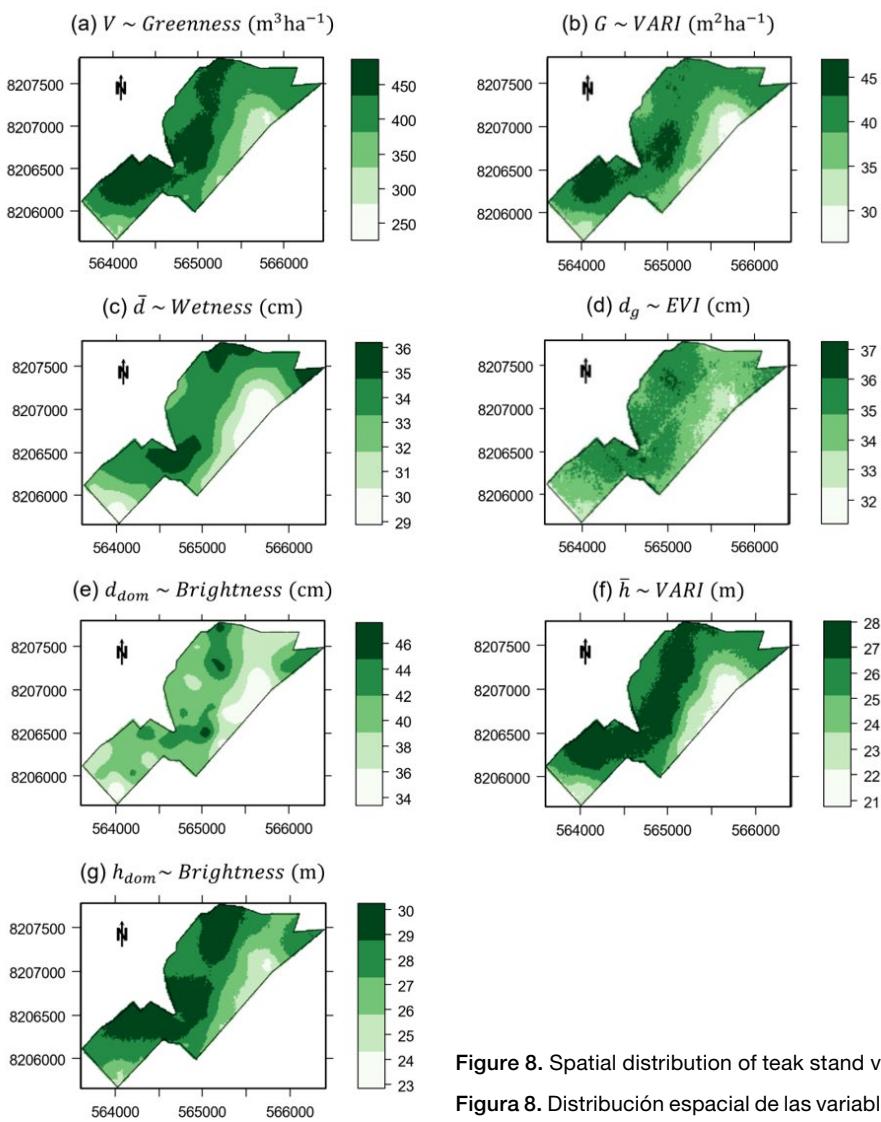


Figure 8. Spatial distribution of teak stand variables by kriging with external drift.

Figura 8. Distribución espacial de las variables del rodal de teca mediante kriging con deriva externa.

**Table 2.** Cross-validation statistics of kriging methods for estimating stand variables.**Cuadro 2.** Estadísticas de validación cruzada de kriging para la estimación de las variables del rodal.

Method	Variable	Covariable	BIAS	MAE	MPE%	RMSE	RMSE%	$r_{xx}$
Ordinary kriging	V		-0.0035	0.139	2.32	0.206	3.46	0.531
Kriging with external drift	V	Greenness	0.0048	0.148	2.48	0.228	3.83	0.430
Ordinary kriging	G		-0.0027	0.100	2.73	0.148	4.03	0.499
Kriging with external drift	G	VARI	-0.0028	0.104	2.84	0.150	4.09	0.480
Ordinary kriging	$\bar{d}$		-0.0002	0.056	1.60	0.071	2.05	0.539
Kriging with external drift	$\bar{d}$	Wetness	0.0013	0.057	1.65	0.073	2.11	0.501
Ordinary kriging	$d_g$		-0.0007	0.067	1.90	0.090	2.56	0.329
Kriging with external drift	$d_g$	EVI	-0.0005	0.067	1.90	0.089	2.51	0.379
Ordinary kriging	$d_{dom}$		-0.0006	0.0644	1.75	0.087	2.38	0.317
Kriging with external drift	$d_{dom}$	Brightness	-0.0002	0.0655	1.78	0.088	2.40	0.292
Ordinary kriging	$\bar{h}$		-0.0011	0.053	1.64	0.076	2.34	0.629
Kriging with external drift	$\bar{h}$	VARI	0.0013	0.058	1.75	0.076	2.38	0.613
Ordinary kriging	$h_{dom}$		-0.0009	0.058	1.74	0.076	2.3	0.512
Kriging with external drift	$h_{dom}$	Brightness	0.0007	0.061	1.83	0.080	2.41	0.450

The cross-validation statistics by the external drift kriging method presented better results for the  $d_g$  variable when compared to ordinary kriging (Table 2). Although the estimates of forest variables have not, in general, presented higher cross-validation statistics than ordinary kriging, through kriging with external drift it was possible to compose thematic maps with a higher level of spatial detail (Figure 8), ensuring lower average errors. These results are desirable when the objective is to apply precision management in commercial stands since they also allow for adequate estimates in unsampled locations [14].

## Conclusion

Remote sensing variables show a weak linear correlation with teak variables, which tends to make them unviable to use as covariates in geostatistical modeling. However, kriging with external drift predicts spatial patterns of forest variables with greater detail, which results in lower possible smoothing errors than those obtained by ordinary kriging and provides more accurate recommendations for localized management in teak stand.

The integration of remote sensing variables in forest inventory through geostatistics is advantageous since the product of kriging with external drift estimates can be used as auxiliary information. Additionally, as a recommendation, remote sensing variables make it possible to increase the reliability of forest inventories without increasing costs with higher sample intensity.

## References

- [1] J. C. Deb, S. Phinn, N. Butt, C. A. Mcalpine, J. C. Deb, e N. Butt, "Climatic-induced shifts in the distribution of teak (*Tectona grandis*) in Tropical Asia : Implications for forest management and planning", Environ. Manage., p. 422–435, 2017, doi: 10.1007/s00267-017-0884-6.
- [2] V. P. Tewari e B. Singh, "A first-approximation simple dynamic growth model for forest teak plantations in Gujarat state of india", South. For., vol. 80, no 1, p. 59–65, 2017, doi: 10.2989/20702620.2016.1277644.
- [3] C. A. M. Passos, B. L. Junior, e M. R. Gonçalves, "Avaliação silvicultural de *Tectona grandis* L.f., em Cáceres – MT, Brasil: resultados preliminares", Ciência Florest., vol. 16, no 2, Apr., p. 225–232, 2006.

- [4] ABRAF, "Anuário Estatístico ABRAF", 2013. [Online], Available: <http://www.bibliotecaflorestal.ufv.br/handle/123456789/3910>
- [5] IBÁ (Brazilian Tree Industry), "Report 2019", São Paulo, 2019. [Online], Available: <https://iba.org/datafiles/publicacoes/relatorios/iba-relatorioanual2019.pdf>
- [6] W. Kollert e M. Kleine, "The Global Teak Study. Analysis, Evaluation and Future Potential of Teak", Resources, vol. 36, no July, 2017.
- [7] R. Drescher, F. Henrique, S. Moraes, A. L. Pelissari, e F. C. Acosta, "Equações volumétricas para *Tectona grandis* Linn F. em povoamentos jovens no estado de Mato Grosso", Adv. For. Sci., vol. 1, no 2, Jun., p. 65–70, 2014.
- [8] R. S. da Silva, D. G. S. Vendruscolo, J. R. M. da Rocha, A. G. S. Chaves, H. S. Souza, e A. S. da Motta, "Desempenho silvicultural de *Tectona grandis* L. f. em diferentes espaçamentos em Cáceres, MT", Floresta e Ambient., vol. 23, no 3, p. 397–405, 2016, doi: 10.1590/2179-8087.143015.
- [9] F. A. Leal, E. A. T. Matricardi, e E. P. Miguel, "Interpolador geoestatístico para estimar volume num povoamento de *Eucalyptus urophylla*, em Rio Verde/Goiás", Nucleus, vol. 11, no 1, p. 153–160, 2014, doi: 10.3738/1982.2278.1021.
- [10] M. Roveda, A. F. Figueiredo Filho, A. L. Pelissari, e R. M. M. dos Santos, "Continuidade espacial do número de árvores e da área basal de uma floresta ombrófila mista sob diferentes formas e tamanhos de unidades amostrais", Cerne, vol. 22, no 2, p. 189–196, 2016, doi: 10.1590/01047760201622022139.
- [11] G. Matheron, The Theory of Regionalized Variables and its Applications. Les Cahiers du Centre de Morphologie Mathématique in Fontainebleau, Paris. 1971.
- [12] M. A. Oliver e R. Webster, "A tutorial guide to geostatistics: Computing and modelling variograms and kriging", Catena, vol. 113, p. 56–69, 2014, doi: 10.1016/j.catena.2013.09.006.
- [13] N G. Karahan e S. Erahin, "Geostatistics in characterizing spatial variability of forest ecosystems", Eurasian J. For. Sci., vol. 6, no 1, Jan., p. 9–22, 2018.
- [14] A. L. Pelissari, M. Roveda, S. F. Caldeira, C. R. Sanquette, A. P. D. Corte, e C. K. Rodrigues, "Geostatistical modeling of timber volume spatial variability for *Tectona grandis* L. f. precision forestry", Cerne, vol. 23, no 1, p. 115–122, 2017, doi: 10.1590/01047760201723012291.
- [15] R. L. Gil, L. R. Lansanova, L. A. Chig, A. T. de A. Garrett, R. V de Oliveira, e Filho P. C. de Oliveira and K. R. Ferreira, "Modelling spatial variability of soil chemical attributes in *Tectona grandis* stands in central-west Brazil", J. Trop. For. Sci., vol. 30, no 3, p. 376–383, 2018, doi: 10.26525/jtfs2018.30.3.376383.
- [16] J. L. Hernández-stefanoni, J. A. Gallardo-cruz, J. A. Meave, e J. M. Dupuy, "Combining geostatistical models and remotely sensed data to improve tropical tree richness mapping", Ecol. Indic., vol. 11, no 5, p. 1046–1056, 2011, doi: 10.1016/j.ecolind.2010.11.003.
- [17] A. A. dos Reis, J. M. De Mello, M. R. Raimundo, F. W. Acerbi Júnior, M. S. de Oliveira, e J. M. F. de S. Diniz, "Estratificação de um povoamento de eucalipto por interpoladores geoestatísticos e sensoriamento remoto", Pesqui. Agropecuária Bras., vol. 5, no 10, p. 1751–1761, 2016, doi: 10.1590/S0100-204X2016001000007.
- [18] F. Zakeri e G. Mariethoz, "A review of geostatistical simulation models applied to satellite remote sensing: Methods and applications", Remote Sens. Environ., vol. 259, no February, p. 112381, 2021, doi: 10.1016/j.rse.2021.112381.
- [19] A. A. dos Reis, F. W. Acerbi Júnior, J. M. de Mello, L. M. T. de Carvalho, e L. R. Gomide, "Relationship between spectral data and dendrometric variables in *Eucalyptus* sp. stands", Floresta e Ambient., vol. 25, no 2, p. 2–9, 2018, doi: 10.1590/2179-8087.017015.
- [20] E. Grabska e J. Socha, "Evaluating the effect of stand properties and site conditions on the forest reflectance from Sentinel-2 time series", PLoS One, vol. 16, no 3, p. 23, 2021, doi: 10.1371/journal.pone.0248459.
- [21] C. T. Pertille, M. F. Nicoletti, L. R. Topanotti, e L. P. B. Schorr, "Prediction of dendrometric variables by optical data in *Pinus taeda* L. in the south of Brazil", Adv. For. Sci., vol. 8, no 2, p. 1433–1443, 2021, doi: 10.34062/afs.v8i2.9875.
- [22] A. M. Lechner, G. M. Foody, e D. S. Boyd, "Applications in remote sensing to forest ecology and management", One Earth, vol. 2, no 5, p. 405–412, 2020, doi: 10.1016/j.oneear.2020.05.001.
- [23] A. A. dos Reis, J. M. de Mello, F. W. A. Júnior, e L. M. T. de Carvalho, "Estratificação em cerrado sensu stricto a partir de imagens de sensoriamento remoto e técnicas geoestatísticas", Sci. For., vol. 43, no 106, Jun., p. 377–386, 2015.
- [24] K. Bzdga, A. Zarychta, A. Urbisz, S. Szporak-Wasilewska, M. Ludynia, B. Fojcik, e B. Tokarska-Guzik, "Geostatistical models with the use of hyperspectral data and seasonal variation – A new approach for evaluating the risk posed by invasive plants", Ecol. Indic., vol. 121, 2021, doi: 10.1016/j.ecolind.2020.107204.
- [25] C. A. Alvares, J. L. Stape, P. C. Sentelhas, L. D. M. Gonçalves, e G. Sparovek, "Köppen's climate classification map for Brazil", Meteorol. Zeitschrift, vol. 22, no 6, p. 711–728, 2014, doi: 10.1127/0941-2948/2013/0507.
- [26] USGS - United States Geology Survey, "Earth Explorer", 2019. [Online]. Available: <https://earthexplorer.usgs.gov/> (Accessed Jul. 04, 2019).
- [27] QGIS Development team. [CD-ROM]. QGIS Geographic Information System. Open Source Geospatial Foundation Project, 2009.
- [28] B. Duchemin, R. Hadria, S. Erraki, e G. Boulet, "Monitoring wheat phenology and irrigation in Central Morocco : On the use of relationships between evapotranspiration , crops coefficients , leaf area index and remotely-sensed vegetation indices", Agric. Water Manag., vol. 79, p. 1–27, 2006, doi: 10.1016/j.agwat.2005.02.013.
- [29] J. R. Jensen e J. C. N. Epiphanius, Sensoriamento remoto do ambiente: uma perspectiva em recursos terrestres. Parêntese Editora São José dos Campos, 2009.
- [30] D. C. Zanotta, M. P. Ferreira, e M. Zortea, Processamento de

- imagens de satélite. Oficina de Textos, 2019.
- [31] J. K. Yamamoto e P. M. B. Landim, Geoestatística: conceitos e aplicações. Oficina de textos, 2015.
- [32] E. J. Pebesma, "Gstat : multivariable geostatistics for S", 2003, p. 2.
- [33] R. C. Team, "R development core team", RA Lang Env. Stat Comput, vol. 55, p. 275–286, 2013.
- [34] D. Bagnara, P. Domingos, e M. Pietto, "Aplicação da krigagem ordinária na modelagem do pH e da dureza da água subterrânea na área central de Passo Fundo – RS", Teor. e Práica na Eng. Civ., vol. 20, Nov., p. 15–22, 2012.
- [35] Wackernagel. H., "Multivariate geostatistics", in Multivariate Geostatistics: An Introduction with Applications, Springer US, 2003, p. 388.
- [36] K. Le Rest, D. Pinaud, P. Monestiez, J. Chadoeuf, e V. Bretagnolle, "Spatial leave-one-out cross-validation for variable selection in the presence of spatial autocorrelation", Global Ecology and Biogeography, vol. 23, no 7. p. 811–820, 2014, doi: 10.1111/geb.12161.
- [37] S. F. Caldeira, D. Leandro, e C. Oliveira, "Desbaste seletivo em povoamentos de *Tectona grandis* com diferentes idades", Acta Amaz., vol. 38, no 2, p. 223–228, 2008.
- [38] J. C. C. Campos e H. G. Leite, Mensuração florestal: perguntas e respostas, Viçosa, Ed. UFV, 2013.
- [39] S. Péllico Netto e A. Behling, "Data transformation in biological assays", Pesqui. Florest. Bras., vol. 41, 2021, doi: 10.4336/2021.pfb.41e201901916.
- [40] G. S. Nogueira, H. G. Leite, J. C. C. Campos, F. H. Takizawa, e L. Couto, "Avaliação de um modelo de distribuição diamétrica ajustado para povoamentos de *Tectona grandis* submetidos a desbaste", Rev. Árvore, vol. 30, no 3, May., p. 377–387, 2006.
- [41] B. Husch, T. W. Beers, e J. A. Kershaw Jr, Forest mensuration. John Wiley & Sons, 2002.
- [42] E. P. Miguel, A. V. Rezende, F. A. Leal, E. Aparecido, e T. Matricardi, "Redes neurais artificiais para a modelagem do volume de madeira e biomassa do cerradão com dados de satélite", Pesqui. Agropecuária Bras., vol. 50, no 9, p. 829–839, 2015, doi: 10.1590/S0100-204X2015000900012.
- [43] V. Bewick, L. Cheek, e J. Ball, "Statistics review 7: Correlation and regression", Crit. Care, vol. 7, p. 451–459, 2003, doi: 10.1186/cc2401.
- [44] C. C. Aggarwal, Outlier Analysis. Cham: Springer International Publishing, 2016.
- [45] J. Xue e B. Su, "Significant remote sensing vegetation indices: A review of developments and applications", J. Sensors, vol. 2017, p. 1–17, 2017, doi: 10.1155/2017/1353691.
- [46] P. Schober, C. Boer, e L. A. Schwarte, "Correlation coefficients: appropriate use and interpretation", vol. 126, no 5, p. 1763–1768, 2018, doi: 10.1213/ANE.0000000000002864.
- [47] D. Gumprecht, W. G. Müller, e J. M. Rodríguez-Díaz, "Designs for detecting spatial dependence", Geogr. Anal., vol. 41, no 2, p. 127–143, 2009, doi: 10.1111/j.1538-4632.2009.00736.x.
- [48] D. Marusig, F. Petruzzellis, M. Tomasella, R. Napolitano, A. Altobelli, e A. Nardini, "Correlation of field-measured and remotely sensed plant water status as a tool to monitor the risk of drought-induced forest decline", Forests, vol. 11, no 1, p. 77, 2020, doi: 10.3390/f11010077.
- [49] S. V. Ollinger, "Sources of variability in canopy reflectance and the convergent properties of plants", New Phytol., vol. 189, no 2, p. 375–394, 2011, doi: 10.1111/j.1469-8137.2010.03536.x.
- [50] J. Liu, E. Pattey, e G. Jégo, "Assessment of vegetation indices for regional crop green LAI estimation from Landsat images over multiple growing seasons", Remote Sens. Environ., vol. 123, p. 347–358, ago. 2012, doi: 10.1016/j.rse.2012.04.002.
- [51] I. Pfutz, A. Pelissari, C. Rodrigues, S. Caldeira, e A. Corte, "Geostatistical modelling of below- and aboveground carbon stocks of teak stands at different ages", J. Trop. For. Sci., vol. 33, no 2, p. 105–112, 2021, doi: 10.26525/jtfs2021.33.2.105.
- [52] R. B. Gramacy e H. K. H. Lee, "Cases for the nugget in modeling computer experiments", Stat. Comput., vol. 22, no 3, p. 713–722, 2012, doi: 10.1007/s11222-010-9224-x.